Coherent Inference on Optimal Play in Game Trees

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Abstract
Round-based games are an instance of discrete planning problems. Some of the best contemporary game tree search algorithms use random roll-outs as data. Relying on a good policy, they learn on-policy values by propagating information upwards in the tree, but not between sibling nodes. Here, we present a generative model and a corresponding approximate message passing scheme for inference on the optimal, off-policy value of nodes in smooth AND/OR trees, given random roll-outs. The crucial insight is that the distribution of values in game trees is not completely arbitrary. We define a generative model of the on-policy values using a latent score for each state, representing the value under the random roll-out policy. Inference on the values under the optimal policy separates into an inductive, pre-data step and a deductive, post-data part. Both can be solved approximately with Expectation Propagation, allowing off-policy value inference for any node in the (exponentially big) tree in linear time.

1 Introduction
Many round-based two-player games, like Chess and Go, can be represented, up to transpositions, by graphs with the structure of AND/OR trees (Nilsson, 1971). If the branching factor $b$ is constant, a tree of depth $d$ contains $b^d$ nodes. For Go, $b$ and $d$ are on the order of 200, so it seems finding the optimal path through the game should be intractable. Yet humans can find good paths through the Go tree in finite time.

A crucial property of games that humans use is that the tree has structure: Winning positions are not randomly distributed among the tree’s leaves, but clustered. This structure can be modeled by a latent score, representing the amount by which one player is ‘ahead’ or ‘behind’. Random play leads to a random walk, typically changing the evaluation by small increments. Critical moves, changing the score drastically, are a rare occurrence in the tree overall.

Although it is not usually mentioned explicitly, this smoothness is a crucial intuition behind Monte Carlo algorithms for ‘best first’ tree search, like UCT (Kocsis and Szepesvári, 2006), which have been very successful recently. These algorithms repeatedly play roll-outs—random descents through the tree to a leaf, generated by a relatively weak or uniformly random policy. The search tree is expanded asymmetrically from the root, based on the frequency of wins and losses in the roll-outs. If wins and losses were distributed uniformly at random among the leaves, the roll-out results would be almost completely uninformative (Pearl, 1985). The best contemporary Go machines are using UCT as part of their method (Gelly and Silver, 2008). However, UCT-like methods base their value estimates directly on average outcomes of tree-descents, making them dependent on a good exploration and roll-out policy. They also do not propagate information laterally in the tree, although, thanks to the smoothness of the tree, the value of a node does contain information about the value of its siblings.

In this paper, we construct an explicit generative model for the value of game tree nodes under the random roll-out policy. Finding the value under the optimal policy would amount to solving a min-max optimization problem of complexity $O(b^d)$ if all nodes were observed. However, for best-first search algorithms, we will show how an approximate closed form for the unobserved parts of the tree can be constructed and used to derive an approximate message passing scheme. The resulting algorithm tracks a joint posterior belief over the optimal values of all nodes in the
tree. It incorporates a new roll-out at depth $k$ from
the root in $O(kb)$ time (using heuristics, this can be
brought to $O(k)$, the complexity class of UCT), ar-
riving at an intermediate set of local marginals which
is sufficient to evaluate the posterior of an arbitrary
node at depth $\ell$ in the tree (something classic Monte
Carlo algorithms cannot do) with cost $O(\ell)$. The
algorithm can be interpreted as an instance of Bayesian
off-policy reinforcement learning (Watkins and Dayan,
1992; Dearden et al., 1998), inferring the optimal pol-
icy from samples generated by a non-optimal policy.
Our method might be applicable, to varying degree,
to other tree-structured optimization problems, if they
exhibit a functional relationship between steps; c.f.
the metric (though not tree-structured) sets of band-
its considered by Kleinberg et al. (2008).

Our main contributions are the generative model
for the score of game positions, the formulation of a prob-
abilistic best-first tree search algorithm, and a dem-
stration of Expectation Propagation on min-max trees
(the necessary mathematical derivations are part of a
technical report by one of the authors (Hennig, 2009)).

Probabilistic approaches to game tree search have been
suggested before (see e.g. Baum and Smith, 1997; Rus-
sell and Wefald, 1991; Palay, 1985; Stern et al., 2007).
These works concentrated on guiding the search policy.
Here we focus on efficient and consistent off-policy in-
ference for the entire tree. This is valuable because
a coherent posterior over off-policy values provides
meaningful probabilistic training data for Bayesian al-
gorithms attempting to learn a generalizing evaluation
function based on features of the game state.

2 Methods

This section defines the problem (2.1), then develops
the algorithm in several steps. We define the gener-
ative model of on-policy and off-policy values (2.2),
show how to perform inference in this model by way
of example (2.3), and finally combine the results into
an explicit algorithm (2.4).

2.1 Problem Definition

We consider a tree-structured graph defining a round-
based, loop-free, zero-sum game between two opposing
players, MAX and MIN with binary outcomes 1 and −1
(“win” and “loss” from MAX’s point of view). MAX is
defined to be the first player to move.

The task is to predict the outcome of the game, assum-
ing optimal play by both players, from any position in

\[ v_i = \begin{cases} 
\max_c \{v_c\} & \text{if \, \text{MAX} \, \text{plays \, at} \, i \\
\min_c \{v_c\} & \text{if \, \text{MIN} \, \text{plays \, at} \, i}
\end{cases} \]

We note in passing that OR-trees (i.e. tree-structured
optimization problems, and non-adversarial games like
Solitaire) are a trivial variant of this formulation.

Figure 1 shows the full generative model for a small
tree of $b = 2$ and $d = 4$, as a directed graphical model.
Note that only the sign of outcomes is observed (black
nodes), all other variables are latent.

2.2 Generative Model

The following two definitions jointly form a generative
model $\mathcal{M}$ for the latent scores $\mathcal{G} = \{g_i\}$ and optimal
values $V = \{v_i\}$ of tree nodes $i$.

Definition: The score $g_i$ of node $i$ models the value of
$i$ under random play. It is a real number such that

- for any terminal position $t$, $\text{sign}(g_t) = r_t$, where
  $r_t$ is the binary result of the game at $t$. The like-
ilhood for $g_t$ is thus a step function (denoted $\theta$):
  \[ p(g_t | g_i) = \theta(r_t g_i) \] (1)

- the score of child node $c$ of node $i$ is generated
  from $g_i$ by a univariate Gaussian step:
  \[ p(g_c | g_i, \mathcal{M}) = \mathcal{N}(g_c; g_i, 1) \] (2)

- the prior for the score of the root note is Gaussian
  \[ p(g_0) = \mathcal{N}(\mu_0, \sigma_0^2) \] (one is free to choose $\mu_0$ and
  $\sigma_0$, although $\mu_0 = 0$ is the obvious choice).

Thus, scores of sibling nodes are independent given
their parent’s score, and roll-outs are Brownian ran-
dom walks. For binary results, the scale factor of the
steps is arbitrary and set to 1 for simplicity only.

Definition: The off-policy value $v_i$ of node $i$ is the
value of node $i$ under optimal play by both players.
That is

- for terminal positions $t$, we have $v_t = g_t$.

- for non-terminal positions $i$ with children \{c\}
  \[ v_i = \begin{cases} 
\max_c \{v_c\} & \text{if \, \text{MAX} \, \text{plays \, at} \, i \\
\min_c \{v_c\} & \text{if \, \text{MIN} \, \text{plays \, at} \, i}
\end{cases} \]

Games with real-valued outcomes are in fact an easier
variant of this problem, because the scores $g_i$ of term-
nal nodes can be observed directly, rather than just their signs.
2.3 Inference

We will derive an approximate Belief Propagation algorithm (Lauritzen and Spiegelhalter, 1988; Pearl, 1988) to perform inference both on the scores and values, using Gaussian Expectation Propagation (Minka, 2001) to project the messages to the normal exponential family. Belief Propagation is a message passing scheme between the nodes of a Bayesian network: The marginal probability \( p(x_i) \) of a latent variable \( x_i \) in the model is the product of messages \( m_{x_j}(x_i) \) from neighboring variables \( x_j \), as defined by the structure of the graph. Gaussian Expectation Propagation keeps inference tractable by replacing these messages with approximate Gaussian messages, such that the resulting marginal matches the first two moments of the true marginal.

Inspecting Figure 1, it might seem like inference in the tree would call for messages among all \( b^d \) nodes. In this Section, we will show that this can be avoided, because the messages from unobserved parts of the tree can be derived \textit{a priori}, in jointly \( O(bd) \) time.

We assume the learner acquires data in a best-first manner: At any given point in time, it tracks an asymmetric but contiguous tree in memory which includes the root. Additional nodes are added to the boundary of the stored tree by requesting a roll-out from that node. The message passing is performed in parallel with the search process. The resulting message passing schedule is quite complex. For clarity, we will use an example descent through Figure 1.

2.3.1 On-Policy Inference on \( g \)

Each search descent begins at the root node 0. As the descent passes through the tracked part of the tree, a policy \( \pi \) chooses among available children, potentially based on the current beliefs over their \( v \). For our example, say the policy chose the descent 0 \( \rightarrow \) 1 \( \rightarrow \) 12. At each step, we update the message from the current node to the chosen child. The message out of \( g_1 \) in the direction of \( g_{12} \) is

\[
m_{pa(1)}(g_1) = \prod_{j \in ch(1)\setminus 12} m_{g_j}(g_1) \equiv \mathcal{N}(\mu_{1\setminus 12}, \sigma_{1\setminus 12}^2) \quad (4)
\]

where \( ch(1)\setminus 12 \) is the set of child nodes of 1 excluding node 12, and \( pa(1) \) is the parent node (i.e. 0) of 1. And the message into \( g_{12} \) is

\[
m_{g_1}(g_{12}) = \int p(g_{12}|g_1)p(g_1|g_{pa(1)}, \{g_{ch(1)\setminus 12}\}) dg_1 = \mathcal{N}(\mu_{1\setminus 12}, \sigma_{1\setminus 12}^2 + 1) \quad (5)
\]

Assume node 12 just reached is not part of the stored tree yet. To add it to the tree, we request a roll-out
starting from 12, which turns out to be of length $\ell = 2$ and have result $r = +1$ (see Figure 1). The data thus gained is a likelihood $p(r|g_i)$ of the score of the terminal position $t$ at the end of the roll-out, which is a step-function and thus not normalizable. However, we can generate a prior over $g_i$ as a message from the current marginal over $g_{12}$ by integrating out the $t$ intermediate steps (see Equation (5)):

$$p(g_i|g_{12}) = N(\mu_{12}, \sigma_{12}^2 + \ell)$$

(6)

giving a posterior over $g_i$, which is a truncated Gaussian. To get the EP message back to $g_{12}$, we need a function $f_{g_{12}}^{EP}$, which calculates the moments of this truncated Gaussian distribution. The calculation is straightforward, because for Gaussians in general

$$\int_0^\infty xN(x; \mu, \sigma^2) = \mu \Phi \left( \frac{\mu}{\sigma} \right) + \sigma \phi \left( \frac{\mu}{\sigma} \right)$$

and

$$\int_0^\infty x^2N(x; \mu, \sigma^2) = \mu^2 \Phi \left( \frac{\mu}{\sigma} \right) + \mu \sigma \phi \left( \frac{\mu}{\sigma} \right) + \sigma^2 \Phi \left( \frac{\mu}{\sigma} \right)$$

This result was used previously for EP by Herbrich et al. (2007). To finally arrive at the message $m_{g_{12}}(g_{12})$ from the roll-out to $g_{12}$, we need to apply $f_{g_{12}}^{EP}$ again to the resulting EP message. Note that the message contains no $g$ other than $g_i$. It is thus possible to perform inference on $g_i$ using exactly two messages: One from its parent node, and one from the outcome of the roll-out.

To incorporate the new knowledge from this roll-out into all parent nodes of $g_{12}$, we pass messages of analogous form to Equation (5) back up the tree. This obviously does not propagate the information through the whole tree, but it leads to a situation where the score $g_i$ of any node $i$ in the tree can be evaluated in linear time, simply by performing one descent from the root towards $i$, updating messages analogously to Equation (5) downwards during the descent.

There is one more pitfall to avoid: The next time a search descent passes through node 12 which currently lies at the boundary of the tracked tree, leading to the addition of a child node of 12 and a roll-out from there, it becomes necessary to remove the information gained from the roll-out at node 12 from the marginal. Because the roll-out necessarily passed through one of $i$’s children, information from that child would otherwise be counted twice. This removal corresponds to ‘dividing’ the corresponding message out of the marginal, which for Gaussians has closed form:

$$N(x; \mu_a, \sigma_a^2)/N(x; \mu_b, \sigma_b^2) \propto N \left( x; \frac{\mu_a^2 - \mu_b}{\sigma_a^2 - \sigma_b^2}, \frac{\sigma_a^2 - \sigma_b^2}{\sigma_a^2 - \sigma_b^2} \right)$$

This ‘division’ operation for distributions is used extensively in EP message passing schemes because it increases computational efficiency (Minka, 2001).

### 2.3.2 Off-Policy Inference on $v$

The previous paragraph sketched the message passing inference leading to a consistent posterior over scores $G$. How can these beliefs be used to obtain a posterior over the values $V$? We will use the definitions of Section 2.2 to again derive a message-passing scheme running parallel to the search process.

First, consider again node $i = 12$ in Figure 1, at the boundary of the stored tree. The value under optimal play is the sum of the two independent variables

$$v_i = g_i + \Delta_i$$

(7)

where $\Delta_i$ is the optimal reachable increment to the score of $i$. So the inference breaks up into a deductive part (on $g_i$, as solved in the previous section) and an inductive part (on $\Delta_i$). If the next move after $i$ is controlled by $\text{MAX}$ (replace max with min in the opposite case), then

$$\Delta_i = \max_{j \in \text{children}(i)} \{\xi_j + \Delta_j\}$$

(8)

where $\xi_j \sim N(0,1)$ is the unknown Brownian step to the score of node $j$. Deriving a belief over $\Delta_i$ for any node $i$, which is $\ell_i$ steps from a terminal position, is a recursive problem which depends only on $\ell_i$, and the branching factor $b$: Assuming $\text{MIN}$ gets the last move (with straightforward variations in other cases), $\Delta_i$ is

$$\Delta_i(\ell_i, b) = \begin{cases} 
\delta(\Delta_i - 0) \\
\max_{j=1...b} \{\Delta_j(\ell_i - 1, b) + \xi_j\} \\
\min_{j=1...b} \{\Delta_j(\ell_i - 1, b) + \xi_j\}
\end{cases}$$

(9)

for $\ell = 0$, for $\ell \mod 2 = 0$ and for $\ell \mod 2 = 1$, respectively.

Similarly to the situation in the previous Section, the beliefs generated by this recursive operation are not Gaussian themselves. To perform the EP approximation, we need the function $f_{\text{max/min}}^{EP}$ calculating the moments of the maximum or minimum over Gaussian variables. The necessary derivations are lengthy and thus left out here for brevity. A technical report by one of the authors (Hennig, 2009) develops these moments for the case of the maximum of two Gaussian variables (Equation 25 in the cited work) and shows how to combine such binary comparisons iteratively into an approximate posterior belief over the maximum of a finite set of such variables (Section 3.2 in the cited work). The corresponding messages for the minimum of variables is a trivial variant, because $\min_i \{x_i\} = -\max_i \{-x_i\}$. Using this approximation, we arrive at a recursive operation in closed form, which can be used to derive the message from $\Delta(\ell)$ for all $\ell$ up to a pre-defined depth. This can be done prior
to data acquisition, once for the entire game tree, in $O(bd)$ time, which is easily tractable even for massive game trees like that of $19 \times 19$ Go. In our simple, non-optimized implementation, this step takes about 2 minutes on a contemporary desktop machine, for a tree of Go-like dimensions ($10^{400}$ nodes). This step is a parametrized version of the Monte Carlo technique known as density evolution (e.g. MacKay, 2003, Sec. 47.5). See Section 3.3 for an experimental analysis of the quality of the approximation.

We sum the independent variables $g_i$ and $\Delta_i$, using the exact function

$$f_{\sum_{N}}[\mathcal{N}(\mu_a, \sigma_a^2), \mathcal{N}(\mu_b, \sigma_b^2)] = \mathcal{N}(\mu_a + \mu_b, \sigma_a^2 + \sigma_b^2)$$

For a node $j$ which does not lie on the boundary of the stored tree, $v_j$ is given by Equation (3). Marginals $p(v_c,j)$ are available for all children $c_j$ of $j$ in this case. Hence, an approximate Gaussian message to $v_j$ from its children can be found using the same method as above. However, it is important to note that the children of $v_j$ are correlated variables, because they are all of the form shown in Equation (7), sharing the contribution $g_j$. The EP equations derived in (Hennig, 2009) include the correlated case. Using approximate Gaussian messages

$$q(v_k) = \mathcal{N}(\mu_k, \sigma_k^2) = \mathcal{N}(\mu_{g_j} + \mu_{\Delta_k}, \sigma_{g_j}^2 + \sigma_{\Delta_k}^2)$$

for each child $k$ of $j$, the correlation coefficient $\rho_{k_1k_2}$ between two children $k_1$ and $k_2$ is

$$\rho_{12} = V^{-1}(\sigma_{g_1}^2 - \mu_{g_1}\mu_{\Delta_{k_1}} - \mu_{g_2}\mu_{\Delta_{k_2}} - \mu_{\Delta_{k_1}}\mu_{\Delta_{k_2}})$$

$$V \equiv (\sigma_{g_1}^2 + \sigma_{\Delta_{k_1}}^2 - 2\mu_{g_1}\mu_{\Delta_{k_1}})^{1/2}$$

$$\cdot (\sigma_{g_2}^2 + \sigma_{\Delta_{k_2}}^2 - 2\mu_{g_2}\mu_{\Delta_{k_2}})^{1/2}$$

### 2.3.3 Loopiness

From Figure 1, it is clear that the graphical model contains loops. However, because we assume no outside information on the optimal value $v_0$ of the root, the EP messages from the $v$’s to the $g$’s are uniform (Hennig, 2009). So inference on the $g$’s is not influenced by the $v$-part of the graph. Thus, while the inclusion of the $v$-nodes does create loops in the overall graph, the message-passing within the $v$-part of the tree is based on a consistent set of marginals on the $g$’s.

### 2.4 Algorithm

Algorithm 1 sums up the message-passing and search scheme presented in the previous Sections. It defines a recursive function that descends through the tracked tree to a leaf, passing messages downward (the operator $f_{\mathcal{N}(0,1)}(p)$ refers to Equation (5)). At the boundary of the stored tree, it performs a roll-out, then passes $g$ and $v$ messages upwards to the root. The actual top-level search algorithm repeatedly calls this function, accumulating more and more data, at roughly constant computational cost per call (apart from the small increase in cost caused by the growth of the stored tree). The descent through the stored tree is guided by some policy $\pi$, which may (and should) depend on the current beliefs over values $v$. The notation $p(a) \land s(i)$ refers to the parent of $i$ and the set of siblings of (and including) $i$, respectively. The function STORED accesses a one-dimensional array of stored inductive messages from the un-explored parts of the tree, as discussed in Section 2.3.2.

#### Algorithm 1 Bayesian Best-First Tree Search

1: procedure DESCENT($i$)
2: if $i$ previously visited
3: ▷ choose child $c$ to explore
4: $c \leftarrow \pi(i)$
5: ▷ update message to $c$.
6: $p(g_c) \leftarrow p(g_c) / m_{g_c}(g_c)$
7: $m_{g_c}(g_c) \leftarrow f_{\mathcal{N}(0,1)}[p(g_c) / m_{g_c}(g_c)]$
8: $p(g_i) \leftarrow p(g_i) \cdot m_{g_c}(g_i)$
9: ▷ continue descent (returns $g$ message from child)
10: $m'_c(g_i) \leftarrow$ DESCENT($c$)
11: ▷ update marginals
12: $p(g_i) \leftarrow p(g_i) / m'_c(g_i)$
13: $m_c(g_i) \leftarrow m'_c(g_i)$
14: else
15: ▷ divide out roll-out from parent’s marginal
16: $p(g_{pa(i)}) \leftarrow p(g_{pa(i)}) / m_{g_{pa(i)}}(g_{pa(i)})$
17: $m_{pa}(g_{pa(i)}) \leftarrow \mathcal{N}(0, \infty)$
18: ▷ update message from parent to $g_i$ (lines 3 to 5)
19: ▷ do roll-out
20: $(r_i, \ell_i) \leftarrow$ ROLLOUT($i$)
21: ▷ generate message from roll-out result to $i$
22: $m_{r_i}(g_i) \leftarrow f_{\mathcal{N}(0,\ell_i)}[\mathcal{N}_{\text{EP}}(r_i, f_{\mathcal{N}(0,\ell_i)}[p(g_i)])]
23: $p(g_i) \leftarrow p(g_i) \cdot m_{r_i}(g_i)$
24: ▷ generate marginal for $v_i$
25: $p(v_i) \leftarrow f_{\sum_{N}}[p(g_i), \text{STORED}(\ell_i)]$
26: end if
27: ▷ Calculate messages to parent’s $g$ and $v$
28: $m_i(g_{pa(i)}) \leftarrow f_{\mathcal{N}(0,1)}(p(g_i) / m_{pa}(g_{pa(i)})$
29: $p(v_{pa(i)}) \leftarrow f_{\mathcal{N}(0,1)}(\min\{p(v_k) \land s(i), p(g_i)\})$
30: return $m_i(g_{pa(i)})$
31: end procedure

### 2.5 Exploration Policies

While the inference process itself is independent of the chosen policy $\pi$ in Algorithm 1, the policy is crucial to make the roll-outs informative about the optimal path (this is a general feature of off-policy reinforcement learning). Many possible policies are available,
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Figure 2: Log likelihood of the ground truth value of $g_i$ (crosses) and $v_i$ (triangles) under the beliefs of the model, at varying depth from the root in artificial trees. Averages over all nodes at those depths.

among them the point-estimate based UCT (Kocsis and Szepesvári, 2006), greedy choice among samples (Thompson, 1933), information gain (Dearden et al., 1998) and ‘optimistic’ exploration based on weighted sums of mean and variance of the value estimates. The latter approach has received some interest recently (Kolter and Ng, 2009), but the ‘right’ choice of policy is still a matter of open debate. In our experiments, we used optimistic exploration where applicable.

3 Results

We performed experiments to answer several questions arising with regard to the described inference model: Is the model of Brownian motion applicable for real games, like Go (3.1)? Is the EP approximation on the roll-out results effective (3.2)? Does the recursive min/max approximation in the inductive part of the inference produce a reasonable approximation of the true min/max values (3.3)? How robust is the model to mis-match between generative model and ground truth (3.4)? And could the algorithm be used as a standalone searcher (3.5)?

3.1 Structure of Go Game Trees

We generated one random path through the tree of $9 \times 9$ Go. At each level in this path, roll-outs from all legal moves $i$ at this position were generated (1000 roll-outs from each $i$). Depending on the depth from the root, there were between 81 and 0 such legal positions. We stopped the game after 71 steps, when there were less than 5 legal moves available. The average length $\ell_i$ of the roll-outs and the empirical frequency $\hat{p}(\text{win}|i)$ of a win for the MAX player from $i$ under a random roll-out policy was stored. This implicitly defines the value of $g_i$ under the model through

$$p(\text{win}|g_i) = \int_{g_i}^{\infty} N(g_i, g_i, \ell_i) \, dg_i = \Phi \left( \frac{g_i}{\sqrt{\ell_i}} \right)$$

(11)

where we replace $p(\text{win}|g_i)$ and $\ell_i$ with empirical averages. With sufficiently many samples, $p(\text{win}|g_i)$ and thus $g_i$ can be evaluated up to negligible error. Our model $\mathcal{M}$ then amounts to the statement that the $\{g_i\}$ of sibling nodes in the tree are distributed like the standard normal distribution around their parent’s value.

Figure 3 shows Q-Q plots of these empirical distributions against $\Phi$ for depths $d = 0$ to 71 from the root. A standard normal would lie on the diagonal in this plot; weaker tailed distributions are steeper, heavier tailed ones flatter. The left plot shows results from a uniform roll-out policy (only excluding illegal and trivially bad ‘suicide’ moves). Given the limited sample sizes, especially towards the end of the game, the empirical distributions are strikingly similar to $\Phi(x)$.

Contemporary Monte Carlo search algorithms use ‘heavy roll-outs’, i.e. policies that produce less random, more informative results. Clearly, the hypothesis of ergodicity in our model will be more and more invalid the smarter the policy (the limit of a perfect policy would repeatedly generate only a single perfect roll-out). To examine how drastic this effect is, we repeated the above experiment with a smart policy, similar to the published parts of MoGo (Gelly et al., 2006) (Figure 3, right). While the results do develop heavier tails, the overlap with $\Phi(x)$ is still quite good.

3.2 Inference on the Generators

A good way to evaluate the quality of Bayesian models is the (log) likelihood they assign to ground truth in known test environments. To do so, we generated 500 artificial trees of $b = 2, d = 18$ from $\mathcal{M}$. The inference model was implemented as presented in Algorithm 1. A time step corresponds to one descent into the tree, ending with a roll-out at a previously unexplored node. For the descent through previously visited nodes, an optimistic policy was used (see Section 2.5), choosing greedily among children $i$ based on $\mu_i + 3\sigma_i$. In addition, the value of the root node’s score was assumed to be 0 with high precision. Figure 2 shows the log likelihood assigned to the ground truth of both $g$ and $\nu$ at distances 1, 2 and 3 from the root.

As expected, the likelihoods rise on average during learning. They saturate when the majority of the nodes is expanded, because the (binary) roll-out results do not contain sufficient information to determine $g_i$ and thus $\nu$, to arbitrary precision. Nodes deeper in the tree saturate at smaller likelihoods because they receive information from fewer offspring nodes. They
3.3 Recursive Inductive Inference on $v$

To evaluate the quality of the inductive part of the approximation, 1000 artificial game trees were generated and solved by explicit min/max search as in the preceding section. Figure 4 compares empirical ground truth of $\Delta$ and values predicted by pre-data inference, for all nodes at two different distances $\ell$ from the leaves, in 1000 artificial game trees. Despite the repeated application of the Gaussian approximation to non-Gaussian beliefs, there is good agreement between predictions and ground truth.

3.4 Errors Introduced by Model Mismatch

For the last two experiments, the artificial game trees were generated by the generative model $M$, and we showed in Section 3.1 that the real-world game Go is in fact pretty close to this model. However, other games might be less well approximated by $M$. As a tentative test of the severity of the errors thus introduced, the Bayesian searcher was tested on 500 generated $p$-game trees (Kocsis and Szepesvári, 2006) with $b = 2, d = 18$. These trees are also generated by a latent stochastic variable, but with a different generative model, choosing $\xi_i$ uniformly from $[-1, 0]$ if the player is MIN and uniformly from $[0, 1]$ if the player is MAX.

We performed a best-first search in these trees (results not shown due to lack of space), and compared to the performance on trees for which $M$ is the correct model. The performance on these two types of trees during the search was very similar, except for a globally slightly higher chance of the model misclassifying nodes into winning and losing nodes. At least in this particular case, the model mismatch causes only minor decay in performance.

3.5 Use as a Standalone Tree Search

It is tempting to interpret the presented inference algorithm as a standalone search method. Experiments on artificial game trees (results not shown) suggest that the resulting algorithm is rarely much better at exploring the tree than e.g. a vanilla UCT searcher, and that performance depends strongly on the policy used. The
intent of the presented algorithm is not to develop a good tree searcher, but to provide a consistent posterior from which generalizing evaluation functions can be learned.

4 Conclusion

We have presented a generative model for game trees, and derived an approximate message-passing scheme for inference on the optimal value of game positions, given samples from random roll-outs. Inference separates into two tractable deductive and inductive parts, and allows evaluation of the marginal value of any node in the tree in linear time. Similar to the way humans think about games, the computational cost of our algorithm depends more directly on the number of data collected than on the size of the game tree.

Like any model, the assumption of Brownian motion is imperfect. Neither does it catch all the details of any one game, nor does it necessarily apply to all games. But it provides a quantitative concept of a smoothly developing game between competing players. Our experimental results suggest that errors introduced by model mismatch are not severe, and that this is a strikingly good model of Go. We argue that it lies at a ‘sweet spot’ between unstructured priors and non-probabilistic, rule-based methods, which can over-fit.

Retaining a posterior does not necessarily improve on point-based methods in the search for the optimal next move. Nevertheless, the results presented here are valuable in two ways: humans use sophisticated feature-based concepts to play games, not full tree search. Emulating this behavior with a Bayesian algorithm depends more directly on the number of data collected than on the size of the game tree.

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References


